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Dynamic Connectedness and Hedging Opportunity Nexus between Clean Energy, Crude Oil and Technology Sector

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Abstract

In this paper, dynamic connectedness and time varying hedging opportunities between WilderHill clean energy ETF, West Texas Intermediate (WTI) crude oil and Arca Tech 100 ETFs were analyzed between 3 May 2005-22 October 2021. The volatility interdependency and conditional correlation nexus were investigated by TVP-VAR and DCC-FIGARCH model with daily frequencies. TVP-VAR results prove that dynamic connectedness increases among assets, especially during periods of turbulence such as Covid-19. Furthermore, DCC-GARCH model results show that ETFs included in the analysis exhibit long memory properties. The conditional correlation between ECO and PSE is around 71%. The most important finding of the research is that long position risks arising in both ECO and PSE can be effectively and efficiently hedged with WTI. On the other hand, it was determined that WTI can be added to the portfolio in order to reduce the risks of portfolio to be established with clean energy and technology sector. Another remarkable result of the paper is that the simultaneous evaluation of ECO and PSE in portfolio strategies cannot contribute to risk minimization.

Keywords: Dynamic Connectedness; Time varying hedging opportunities; Clean Energy; WTI; Technology sector

Jel Codes: G11; Q42

Introduction

Supply of energy plays an important role in today's society, ranging from assuring basic human needs to independence of countries. There are three basic sources where can be provided.

Traditional fossil sources like crude oil which has been in use for nearly more than a century, from renewable energy sources and from nuclear raw materials in the form of nuclear energy. However crude oil prices are determined according to demand and supply principle, local and international problems of the crude oil exporting countries which are called "OPEC", sudden shocks in the market, like contraction of demand or political and social restrictions taken for oil and its derivatives due to global climate change will cause high volatility in the price changes. On the other hand, boost of oil price will trigger the demand on alternative sources, of course this will make a positive impact on the revenue stream of such companies. Although renewable energy capacity has doubled globally from 2007 to 2016, crude oil and other liquids share on global energy consumption is still around 32% [1,2]. Although crude oil prices had gone down to 32 \$/ barrel in 2008 crisis, than increased to 114 \$/ barrel in 2011 and then went down to 26 \$/barrel in

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2016, which is a loss of 77 % compared to 2011 prices. Later, on prices differed between 51 / 77 \$ a barrel. During Covid19 pandemic due to decline of demand prices went down to 20 \$/barrel but recovered to 30 \$. Rise in the profits of Tech companies related with clean energy companies are highly expected due to this unstability of oil prices and markets [3].

ARCH and its derivative traditional short memory models are used in most studies investigating the determination of hedging effectiveness and portfolio diversification opportunities. However, in many empirical studies with financial statistics, it has been determined that the autocorrelations of the return and volatility series remain nonzero for a fairly wide delay (using remain non-zero expression of Ding et al. [4], Baillie et al. [5], Ding and Granger [6], Andersen et al. [7]. In all of these studies, it has been proven that autocovariance functions disappear at a slow rate. The most important originality of this study is that the volatility structures of the series, in which portfolio optimization and hedging opportunities are investigated, used the FIGARCH model, which takes into account the long memory (GARCH, EGARCH, APARCH, etc.) instead of short memory models in multivariate form.

The rapid progress in the renewable energy and technology sectors in recent years has reached remarkable levels. According to the "Global Trends in Renewable Energy Investment 2020" report of the UN Environment Program (2020), the investment made in renewable energy in the 2010-2019 period reached 2.7 trillion dollars. Although the Covid-19 process is delayed, it is planned to make an additional \$1 trillion additional non-hydro renewable energy investment until 2030 [8]. Recently, modeling and forecast of the volatility of financial assets with robust methods attracts the attention of investors, especially in portfolio diversification. This paper basically has two different aims. Firstly, this study aims to investigate the dynamic connectedness and volatility transmission channels between oil, clean energy and technology ETFs. This is one of the most important motivations of the paper, which is thought to fill the gap in the literature. We employ a new approach to the Diebold and Yilmaz [9] connectedness index, developed by Antonakakis and Gabauer [10], which is based on TVP-VAR. Secondly, this study is to demonstrate the conditional correlations between clean energy, technology and wti, as well as to measure the hedging opportunities of long position risks arising from investments made in both clean energy and technology sectors and fossil fuels. The second important contribution of the study to the literature is that the DCC-FIGARCH-t method is used in the search of hedging opportunities, unlike the studies in the previous literature.

In the following sections of the paper, firstly, a summary of the studies in the literature is presented, and then the econometric methods used are introduced. In the fourth part, data and the obtained empirical findings are given and the last part includes results and discussions.

Literature Review

There are not much study analyzing relation amidst share values of crude oil companies and alternate energy source and technology companies. The very first one was performed in 2008 by Henriques and Sadorsky [11]. The empirical relation amidst share values of alternate energy source and technology companies and crude oil manufacturing companies were found to have "granger" effect.

Kumar et al. [12] has claimed that alterations in the alternative energy source index is related with crude oil cost and share value of alternative energy source and technology companies, as well with previous alterations in rate of interests. Any rise in crude oil prices affects alternate energy source indices positively. In another study, Sadorsky [13] had performed one of the basic studies about this subject and analyzed the spread of unpredictability amidst crude oil prices and share value of alternate energy source and technology corporations. The results were showing a strong link in high technology share values with alternate energy source company shares values compared with crude oil company shares. If you buy a 20-cent oil share for short term, you can secure this investment with a 1 \$ high technology company share for long term.

Managi and Okimoto [14] analyzed structural breaks in the long run relation of alternate energy source shares and found a positive relation in crude oil and alternate energy source prices after the structural break in 2007. Bondia et al. [15] has found long term relation in one or two endogenous breaking points between oil prices, alternative energy and high-tech company stocks. In addition to this, while alternate energy source and high technology company share values were affected by crude oil prices and interest rates in the short terms but not in long terms.

Zhang and Du [16] showed that alternate energy source company share values have more correlation with high technology company share values rather than crude oil and coal prices. Reboredoa and Ugolini [17] evaluated the effect of cardinality of clean energy share profits in price alterations of fossil fuels (oil, natural gas, coal) and power generating costs. They have found that, whenever there is an up/down fluctuation in power generating costs, it has a major effect on renewable energy price dynamics. Moreover, electric prices in Europe and crude oil prices in United States are major determinants in renewable energy share fluctuations.

Ferrer et al. [18] shows that correlation among these occur in short term, such as up to 5 days, but long-term effects were small in United States. Also, another important result of this study was, neither in long term nor short term crude



oil price has major effect in the performance of alternate source energy corporate shares in the stock exchange market. In 2018 Lee and Baek [19] have used ARDL model which considers asymmetrical effects and nonlinear. It was found that, alterations in crude oil prices have asymmetrical and positive effect on alternate energy source company shares in short term. In a study carried out by Song et al. [20] examining the dynamic directional information spillover of return and volatility between fossil fuel energy market, investor sentiment, and renewable energy stock prices, it was determined that the spread of volatility was stronger than the spread of returns and it suggests that the risk transfer between the markets was more remarkable. The effect of fossil fuel energy markets (especially the crude oil) on the alternative energy companies' shares in the stock exchange markets was higher than that of the sentiments of investors. Finally, investor sentiment towards alternative energy markets can partially explain the profits of these shares and their fluctuations. Magyereh et al. [21] with a different approach from their previous study, examined correlations between crude oil shares and alternate energy source and technology shares. When resolving statistics, it was found that, short term profits from crude oil market shares does not affect and get effected from the profits of alternate source energy and technology shares, but in the long term, there is remarkable transfer of profit as an investment from crude oil shares to alternate source energy and technology corporate shares. Over all scales a strong return link was observed amongst alternate source energy shares and such high technology providing corporate shares. The spread of unpredictability was significant in all statistics and alterations.

Nasreen et al. [3] dynamics of relevance among crude oil profits and alternate source energy and technology corporate share indexes were examined. Obtained findings showed that alternate source energy and technology corporate share indexes are perfect hedging tools for the risks in crude oil market. Examining the oil price and alternative energy company stock portfolio and oil price and technology company stock portfolio, it can be seen that the optimum portfolio is the oil-weighted one. Finally, the authors stated that there was a statistically significant relationship between the crude oil prices and the alternative energy and technology indexes between the years 2006 and 2009.

There are few studies using Diebold and Yilmaz's [22-24] spillover approach. Cronin [25] used this approach to study the relationship between US monetary and financial assets since 2000. He found that sizeable spillovers arose during the periods of economic and financial turbulence after the terrorist attacks to the World Trade Center, the post-Lehman Brothers bankruptcy, etc. Duncan and Kabundi [26] investigated the volatility transmission since it is related with four South African asset classes, namely bonds, commodities, currencies, and equities. The authors found that there was a high level of systemic risk in South Africa and, furthermore, the risk was predominantly related with the country-specific factors. Yilmaz [27] analyzed 10 major East Asian stock markets in order to examine the behavior of return and volatility spillovers across the region over the period between 1992 and 2009. As stated in his study, East Asian stock markets became more interdependent as captured by the increase in return spillovers in the mid-1990s. Caloia et al. [28] studied the strength and direction of semi-volatility spillovers between Germany, France, the Netherland, Italy, and Spain. They found that, over the period between 2000 and 2016, France and the Netherlands were the net donors, while Italy and Spain were the net receivers of both downside risks and upside opportunities. On the other hand, Germany was a net receiver of upside semi-volatility and a net donor of downside semi-volatility. Ji et al. [29] examined the connectedness by using the return and volatility spillovers across six large cryptocurrencies from 7th August 2015 to 22nd February 2018. According to their findings, Bitcoin was found to lose its dominant role in the evolving cryptocurrency market. All the cryptocurrencies were found to alternate between being transmitters and receivers in the course of time. Mensi et al. [30] studied the risk spillover between MSCI world index, S&P 500 index for the United States, stoxx600 index for Europe, P1DOW index for Asia/Pacific, and the five stock markets located in Greece, Ireland, Portugal, Spain, and Italy. They found asymmetric conditional correlations and evidence of significant risk spillovers between these stock markets. Kumar [31] also studied the nature of returns and volatility spillovers between exchange rates and stock price in India, Brazil, and South Africa and found a bidirectional volatility spillover between stock and foreign exchange markets in these countries. Balcilar and Bekun [32] examined the interconnectedness between the returns of the price of oil and foreign exchange on the selected agricultural commodity prices. The results showed a weak pass-through among the investigated variables in rice, sorghum, price inflation, a nominal effective exchange rate and oil price display, while banana, cocoa, groundnut, maize, soybean, and wheat were net transmitters of spillover. Balcilar et al. [33] examined the return and volatility spillover effects in the S&P 500, crude oil, and gold and found a bidirectional return and volatility spillover among these assets. Antonakakis et al. [34] examined the network topology of UK regional property returns over the period 1973Q4-2014Q4 and found that the transmission of inter-regional property return shocks is an important source of fluctuations in the regional property return.

Combining these with the literature, it can be stated that there is a positive relationship between crude oil price and alternative energy source price and there is a significant rise in the alternate energy source indexes whenever crude oil price goes up. Moreover, there is a causal relationship



between technology shares, crude oil prices, and alternate energy source company shares. On the other hand, the relationship between alternative energy company shares and high technology company shares is more intense than the alternative energy company shares and fossil fuel prices.

Research Methodology

Dynamic connectedness

Our empirical analysis consists of the following two steps. In the first step, we examine the volatility contagion effects among our series, in order to understand the transmission mechanism and spillover effects of volatility shocks. In the second section of paper, we estimate the time-varying correlations between WTI, ECO and PSE, as well as, to use this empirical information for the construction of the optimal diversification and hedging opportunity strategies.

The contagion effect of the crises in the financial markets makes it valuable to investigate the dynamic connectedness relationships between assets. Although many methods have been applied to investigate the volatility spillover relationships between markets, the most remarkable paper was developed by Diebold and Yilmaz [22,23,35] who introduced different versions of connectedness procedures based on the notion of forecast error variance decomposition from vector autoregressions [10,36,37].

The research procedure developed by Diebold and Yilmaz [9,22-24,35], especially in the field of economy and finance, volatility spillover, stock market interdependencies, cryptocurrency market contagion etc. has found an extensive field of study on the subject (see Duncan and Kabundi [26]; Yilma [27]; Kumar [31]; Caloia et al. [28]; Antonakakis and Gabauer [10]; Ji et al. [29]; Antonakakis et al. [34]; Balcilar and Bekun [32]; Mensi et al. [30]; Cronin [25]; Balcilar et al. [38].

In order to explore the volatility transmission mechanism in a time-varying form, we use the TVP VAR methodology of Antonakakis and Gabauer [10] that extends the originally proposed dynamic connectedness approach of Diebold and Yılmaz [22,23,35], by allowing the variances to vary via a stochastic volatility Kalman Filter estimation with forgetting factors [39].

According to the Bayesian Information Criterion (BIC) we employ a TVP-VAR (1) with time-varying volatility [10],

 $\Delta x_t = \phi_t \Delta x_{t-1} + \varepsilon_t \qquad \qquad \varepsilon_t \sim N(0, S_t) \tag{1}$

$$\operatorname{vec}(\phi_t) = \operatorname{vec}(\phi_{t-1}) + \xi_t \qquad \qquad \xi_t \sim N(0, \Xi_t) \qquad (2)$$

where Δx_t , ϵt are N × 1 vector and St and Φ_t are N × N dimensional matrices. The parameters $vec(\phi_t)$ and ξ_t are N²×1 dimensional vectors whereas Ξ_t is N²×N² dimensional matrix. Thus, TVP-VAR model can be showed as $x_t = \sum_{i=1}^{p} \phi_{it} x_{t-i} + \varepsilon_t = \sum_{j=1}^{\infty} A_{jt} \varepsilon_{t-j} + \varepsilon_t$. The time-varying

factor of the Vector Moving Average (VMA) model is the important part of the connectedness index served by Diebold and Yilmaz [23] which uses the Generalized Impulse Remponse Function, ψ_{ijt}^{g} , reneralized Forecast Error Variance Decompositions, $\phi_{ijt}^{g}(D)$, formulized by Koop et al. [40] and Pesaran and Shin [41]. We are mainly focused in the GFEVD, which can be formulated by

$$\phi_{ij,t}^{g}(J) = \frac{S_{il,t}^{-1} \sum_{t=1}^{J-1} (t_{t}A_{t}S_{t}t_{j})^{2}}{\sum_{j=1}^{N} \sum_{t=1}^{J-1} (t_{t}A_{t}S_{t}A_{j}t_{i})} \quad \phi_{ij,t}^{g}(J) = \frac{\phi_{ij,t}^{g}(J)}{\sum_{j=1}^{N} \phi_{ij,t}^{g}(J)}$$
(3)

where is a zero vector with unity on the position, $\sum_{j=1}^{N} \tilde{\phi}_{ijt}^{g}(J) = 1$ and $\sum_{i,j=1}^{N} \tilde{\phi}_{ijt}^{g}(J) = N$. where J represents the forecast horizon and it a selection vector with a one on the ith position and zero otherwise. Using the GFEVD, we construct the total connectedness index by

$$C_{i}^{g}(J) = 1 - N^{-1} \sum_{i=1}^{N} \tilde{\phi}_{ii,i}^{g}(J)$$
(4)

Firstly, we are investigate in the spillovers of variables i to all others j, indicating the TDCTO (Total Directional Connectedness To) others defined as;

$$C_{i \to jt}^{g}(J) = \sum_{j=1, i \neq j}^{N} \phi_{ji,t}^{g}(J)$$
(5)

Secondly, we calculate the transmission of all variables j to variable i, called TDCFROM (Total Directional Connentedness From) others formulated as;

$$C^{g}_{i\leftarrow jt}(J) = \sum_{j=1,i\neq j}^{N} \bar{\phi}^{g}_{ij,t}(J)$$
(6)

Lastly, we compute the differencess between TDCTO and TDCFROM in order to gain the NTDC (Net Total Directional Connectedness);

$$C_{i,t}^{g} = C_{i \to j,t}^{g}(J) - C_{i \leftarrow j,t}^{g}(J)$$
(7)

If NTDC is negative, the variable will be interpreted as volatility receiver, if positive, it will be interpreted as volatility transmitter.

DCC-FIGARCH-t Model

MGARCH models are used frequently by researchers to determine portfolio selections, volatility transmission and hedging opportunities between financial markets. Financial series behavior of skewed distributed and leptokurtic character, information shocks eliminate in hyperbolic speed after reaching financial assets, reluctancy of financial series to return to average are causing financial assets to be interpreted as showing long memory behavior. In this respect Fractional GARCH models are preferred instead of GARCH models to examine the volatility structures of financial assets.

In this paper, we will examine dynamic volatility spillover and hedging opportunities among WTI, ECO and PSE. A major advantage of running the DCC-GARCH model developed by Engle [42] is the detection of possible changes in conditional correlations over time; this model allows us

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to detect dynamic investor behavior in response to news and innovations [43].

Engle [42] decomposed conditional covariance matrix as:

$$H_t = D_t R_t D_t \tag{8}$$

$$R_{t} = (diag(Q_{t}))^{-1/2}Q_{t}(diag(Q_{t}))^{-1/2}$$
(9)

$$Q_t = \Omega + \alpha \varepsilon_{t-1} \varepsilon_{t-1} + \beta Q_{t-1} \tag{10}$$

where Rt is defined as the conditional correlation matrix and Dt is <u>a diagonal</u> matrix with time-varying standard deviations $\sqrt{h_{i,t}}$ on the main diagonal. Further, Qt is identified as the approximation of the conditional correlation matrix, defined above in Eq. (8) and (9) as Rt, where the positive semi-definiteness of Qt is guaranteed if both α and β are both positive, while the sum of both α and β is less than one while the initial matrix (Q₁) being positive. $\Omega = (1 - \alpha - \beta)\tilde{R}$, where R representing the unconditional average correlation. We next estimate Dt, which is defined as the conditional volatility through the use of a univariate long-memory (FIGARCH) methodology, where we divide the returns by their conditional volatility and use the $\varepsilon_t = D_t^{-1} R_t$ to estimate the quasi-conditional correlation matrix Q_t. Qt is re-scaled to obtain the conditional correlation matrix described in Eq. (9) [44]. Further, the conditional volatility D and the conditional correlations R_. are then utilized to develop the conditional correlation matrix H_[[45].

The FIGARCH (Fractional Integrated GARCH) model will be annexed to the literature and we will explore the diffusion relationship between long memory dynamics and financial asset returns. We can use the fractionally integrated GARCH (FIGARCH) model of Baillie et al. [5]. This model was generated from the GARCH specification of Bollerslev [46], which is a useful extension of the ARCH process introduced by Engle [47]. The GARCH (p, q) model can be defined as

$$\sigma_t^2 = \omega + \alpha(L)\varepsilon_t^2 + \beta(L)\sigma_t^2 \tag{11}$$

where L is the lag operator; α and β are ARCH and GARCH parameters, respectively; $\alpha(L)\varepsilon_t^2$ and $\beta(L)\sigma_t^2$ offer information about volatility during the previous period and fitted variance from the model during the previous period, respectively; and p and q indicate the order of ARCH and GARCH terms, respectively. Let ε_t be the discrete-time real-valued stochastic process. The conditional variance of FIGARCH (p, d, q) model can be expressed as [48]

$$[1-\beta(L)]\sigma_t^2 = \omega + [1-\beta(L)-\phi(L)(1-L)^d]\varepsilon_t^2$$
(12)

where d denotes the fractional differencing parameter (0 < d < 1) and $(1 - L)^d$ is the fractional differencing operator. The lag polynomials are represented by $\beta(L)$ and $\varphi(L)$. All the roots of $1 - \beta(L)$ and $\varphi(L)$ lie outside the unit circle. FIGARCH (p, \bar{d}, q) model is turned to standard GARCH when $\bar{d} = 0$ and IGARCH model when $\bar{d} = 1$.

Data and Empirical Findings

In this paper we have used data from The WilderHill Clean Energy Index (ECO), NYSE Arca Tech 100 Index (PSE) and daily closing prices of crude oil at West Texas Intermediate (WTI). They are obtained from www.finance.yahoo. The WilderHill Clean Energy is the oldest index, who covers 54 alternate source energy companies. The abbreviation for "Clean Energy Index" in the stock market is "ECO". NYSE Arca Tech 100 Index was founded in 1986 and shows share prices of computer hardware and software companies, health equipment manufacturers, telecommunications and other technology companies. Its abbreviation in the stock market is "PSE" (Table 1).

Dynamic connectedness

We collect daily price of the West Texas Intermediate Crude Oil (WTI), NYSE Arca Tech 100 Index (PSE) and The Wilderhill Clean Energy Index (ECO) between 3 May 2005-22 Oct 2021. In order to examine the Dynamic Connectedness relation, we first calculated the squared return of the series. Figure 1 represents the volatility series pilot. Descriptive statistics of the volatility series is presented in Table 2. We proved that all series are significantly right skewed and all variables are leptokurtic. In addition, JB test results show that, all series distribute non-normal. Furthermore, according to the ERS test results, all series are stationary on the 1% significance level. In addition, the series contain autocorrelation and exhibit ARCH effect.

In Figure 2 and Table 3, we report time-varying dynamic connectedness test results for ECO, WTI and PSE. Although the average Total Connectedness Index (TCI) is 21.56, it can be seen that the dynamic connectedness relation between WTI, ECO and PSE rises between 50 and 65 during some turbulent periods such as Covid-19 (2020), European sovereign debt crisis (2012), Iran sanctions (2012), Lebanon Israel crisis (2006) and Paris climate agreement (2016). Considering the TCI values in table 3, the dynamic connectedness relation between WTI, ECO and PSE ETF series achieved a 21.56. This score provides evidence that there is no strong dynamic connectedness interaction between all variables. For instance, ECO and PSE explain 8.24% and 7.67% of the 10-daysahead forecast error variance of the OIL. Also, ECO explain 20.46% of the forecast error variance of PSE. Similarly, PSE explain 20.76% of the 10-day-ahead forecast error variance of the ECO. In particular, the similarity of the power to explain the forecast error variances of ECO and PSE is similar to the DCC-FIGARCH-t results. This result can be interpreted as the evaluation of ECO and PSE together in portfolio diversification strategies does not contribute to risk minimization.

Figure 3, presents the dynamic net total directional connectedness between clean energy, crude oil and

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Table 1: Information on the data set of the study.

Abbreviation of Variables	Variables Used in the Study	Researches Using the Variables
ECO	The WilderHill Clean Energy Index	Henriques and Sadorsky [11], Kumar et al. [12], Managi and Okimoto [14], Ahmad [49], Reboredoa and Ugolini [17], Ferrer et al. [18], Song et al. [20], Magyereh et al. [21], Nasreen et al. [3]
PSE	NYSE Arca Tech. 100 Index	Henriques and Sadorsky [11], Kumar et al. [12], Managi and Okimoto [14], Bondia et al. [15], Ahmad [49], Ferrer et al. [27], Lee and Baek [19], Nasreen et al. [3]
OIL	West Texas Intermediate Crude Oil	Henriques and Sadorsky [11], [12], Managi and Okimoto [14], Bondia et al. [15], Ahmad [49], Reboredoa and Ugolini [17], Ferrer et. al. [18], Lee and Baek [19], Song et al. [20], Magyereh et al. [21], Nasreen et al. [3]

Table 2: Summary Statistics of Volatility Series.

	OIL	PSE	ECO
Mean	3.629	800.872	1.72
Variance	3550.477	16929508.73	31.453
Channess	46.689***	17.843***	7.059***
Skewness	0	0	0
For Konstantin	2255.81***	456.10***	65.22***
EX.KURTOSIS	0	0	0
10	905002411***	37160057***	790562***
JR	0	0	0
500	-25.598***	-15.747***	-12.814***
ERS	0	0	0
0/40)	964.020***	3604.199***	2586.209***
Q(10)	0	0	0
02/4.02	725.303***	1160.305***	661.533***
Q ² (10)	0	0	0

Notes: *** denote significance at 1%, significance levels respectively; ERS: Stock, Elliott, and Rothenberg (1996) unit-root test; Q(20) and Qs(20) are the empirical statistics of the Ljung-Box test





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Figure 2: Dynamic Total Connectedness based on Eq. (4) of the TVP-VAR (2) model.





		From (j)			
ТО (ј)	OIL	PSE	ECO	From	
OIL	84.09	8.24	7.67	15.91	
PSE	2.57	76.96	20.46	23.04	
ECO	4.99	20.76	74.25	25.75	
ТО	7.57	28.99	28.13	64.69	
Inc. Own	91.66	105.96	102.39	TCI	
NET	-8.34	5.96	2.39	21.56	
Notes: Values represented are variar	ce decompositions for T\/P	VAR(2) model A first-ord	ler lag length was chosen	by Bayesian information	

Table 3: TVP-VAR (2) Total Connectedness.

Notes: Values represented are variance decompositions for TVP-VAR(2) model. A first-order lag length was chosen by Bayesian information criterion.



technology ETFs. We see that OIL (-8.34) is net receiver of the volatility spillover during the may-2005-Oct-2021 period. PSE (5.96) and ECO (2.39) are a net volatility transmitter. The reason for this situation can be explained that there is a volatility slipover from ECO and PSE shares to OIL due to the increasing importance of technology and clean energy companies after the 2000s, and in addition, due to the loss of importance of oil.

Dynamic Conditional Correlation and Hedging Opportunities

In this part of the study, it is aimed to investigate the conditional correlation and hedging opportunities between ECO, WTI and PSE. Firstly, the price series are converted to continuously compounded log returns with $\ln(p_t/p_{(t-1)})$ equation. Summary statistics of log return data series are represented in Table 4. Table show that, all return series ara negatively skewed and leptokurtically distributed. On a final note, the JB test illustrates that the return series of WTI, ECO and PSE exhibit non-normal distribution. In addition, according to the ADF and ERS unit root test results, the return series are stationary. Furthermore, all return series are significantly autocorrelated and exhibit ARCH effects in error.

In Table 5, unconditional correlation parameters show that, positive relationship between PSE and ECO is observed. Although the correlation between both OIL and PSE and OIL & ECO is weak, positive relationship was observed. In Table 6, correlation parameters between squared returns show similarity with the results in Table 5. In Figure 4, volatility clusters are clearly seen in the squared return series.

This result raises doubts about the existence of long memory in asset series. It can be clearly seen from the chart that the 2008 global financial crisis, 2012 European debt crisis, the 2016 Paris climate agreement and the Covid-19 caused a serious increase in the volatility of all 3 series.

In Table 7 shows DCC-FIGARCH-t model results. The estimates of the univariate FIGARCH model (Panel 1) show that the fractionally integrated coefficient "d" is significant for all series. So, this result revealing a high level of shock persistence. d parameter of WTI is higher than the other indices.

In Panel 2 of Table 7 displaying estimation results of DCC. α and β coefficients are positive. Furthermore, β criterion is very close to 1. This reveals a higher persistence of volatility across indices. Also sums of α and β coefficients are <1, indicating estimated DCC criterion scatter in the



Figure 4: Squred Returns (volatility) Plots in The WilderHill Clean Energy Index (ECO), West Texas Intermediate (WTI) crude oil index and NYSE Area Tech 100 Index (PSE).



	OIL	PSE	ECO
Mean	0.0001069	0.0002048	0.0000297
Maximum	0.13882	0.043859	0.068705
Mininum	-0.12256	-0.055313	-0.06791
Std. Dev.	0.011458	0.005577	0.009416
Skewness	-0.36177	-0.413	-0.59015
Excess Kurtosis	14.465	9.8466	7.4081
Jarque-Bera	33893***	15777***	9092.9***
ADF	-36.1956***	-36.3366***	-35.226***
ERS	-28.048***	-27.649***	-29.060***
Q (20)	59.9390***	185.514***	76.7882***
Qs (20)	2731.77***	5007.63***	5756.19***
ARCH (10)	120.84***	159.63***	163.50***

Table 4: Descriptive statistics for daily returns.

Note: Q(20) and Qs(20) are the empirical statistics of the Ljung-Box test for autocorrelation of returns and squared returns series, respectively. ADF refers to the empirical statistics of the Augmented Dickey-Fuller (1979) unit root test respectively. ERS refers to Elliot, Rothenberg and Stock (1996) unit root test. ERS developed a feasible point optimal test, "P-test", which takes serial correlation of the error term into account. The ARCH (10) test proposed by Engle [45] is used to control the validity of ARCH effects. ***implies the rejection of the null hypotheses of normality, unit root, no autocorrelation and conditional homoscedasticity at the 1% significance level.

range of typical GARCH model. The results are showing investment instruments can be used to manage risks arising from another. The diagnostic test results were summarized in Panel 3, DCC-FIGARCH-t model. The Ljung-Box test for the standardized and squared standardized residuals don't reject the null hypothesis of "no serial correlation" for most cases.

Figure 5 shows the conditional correlation between variables. Especially during the Covid 19 period, the conditional correlation between eco and pse increased up to 0.84. Similarly, the correlation between wti and both clean energy and technology indices increased at the beginning of 2020 due to covid-19. Research results are similar to Total connectedness analysis results. The correlation between variables increases during turbulence periods.

Long term positions in ECO, WTI or PSE can hedged with short term positions with other shares. We calculate time varying hedge ratio with the help olf conditional volatility series and be used eq. 13

$$\beta_{ij} = \frac{h_{ij}}{h_{ij}} \tag{13}$$

The conditional volatilities from the DCC-FIGARCH-t model can be used to estimate Time-Varying (TV) hedge ratios. Figure 6 and Table 8 show that, a \$1 long position in OIL can be hedged for 38 cents with a short position in the ECO. Onaverage, a \$1 long position in clean energy companies can be hedged for 27 cents with a short position in the WTI. Furthermore, on average, a \$1 long position in WTI market can be hedged for 43 cents with a short position in the PSE. On the other hand, a \$1 long position in technology stock can be hedged for 13 cents with a short position in the

WTI futures. According to the results, it can be stated that WTI is an inexpensive alternative to manage long position risks arising from both clean energy and technology sectors.

Calculating amount of these assets are important within the optimal portfolios, also calculating short term positions to avoid any long term position risks arising from financial assets. Conditional volatility obtained from DCC-FIGARCH-t can help to calculate amounts of portfolio by using equation 14 and 15.

$$w_{ij,t} = \frac{h_{jj,t} - h_{ij,t}}{h_{ii,t} - 2h_{ij,t} + h_{jj,t}}$$
(14)

$$w_{ij,t} = \begin{cases} 0, if \ w_{ij,t} < 0 \\ w_{ij,t}, \ if \ 0 \le w_{ij,t} \le 1 \\ 1, \ if \ w_{ij,t} > 1 \end{cases}$$
(15)

 $W_{ij,t}$, showing amount of 1st investment in 1\$ investment portfolio, $h_{ij,t}$, showing covariance between these two investments. $h_{jj,t}$, representing variance in both investments. When 1 represents value of asset, the remaining part will show the second investment value in the portfolio. Figure 5 shows time rates of financial asset amounts amongst the prospective portfolios.

The results of the dynamic optimal portfolio weights for two financial asset portfolios comprising of oil and one of the remaining ETFs are represented in Table 9. The mean weight shows the dollar cents that need to be invest in OIL in any \$1 portfolio. For instance, in the OIL/ECO portfolio, 41 cents should be invested in OIL and 59 cents in ECO. Similarly, in the OIL/PSE portfolio, 17 cents should be invested in OIL



and 83 cents in PSE. Due to the high conditional correlation between PSE and ECO, it emerges as a result of including only one of these investment instruments in bilateral portfolios.

Conclusion and Policy Recommendation

Increase in energy demand issue due to energy security arising from the climate changes and economic growth efforts of countries has recently started to take an important place

	OIL	PSE	ECO
OIL	1		
PSE	0.263	1	
ECO	0.309	0.766	1

 Table 6: Unconditional Correlation between daily squared returns.

	OIL _{sqr}	PSE _{sqr}	ECO _{sqr}
OIL	1		
PSE _{sqr}	0.217	1	
ECO _{sqr}	0.267	0.758	1

in the political agendas of countries on a global scale. All these contributed to increased Research and Development works in alternative energy category in the last decade. Especially because of the effects of price shocks caused by uncertainties in oil prices, as in many other industries, the clean energy and technology sector recently drew significant interest in the finance literature. This paper aimed to reveal the time-varying interaction between crude oil and alternative energy & technology industries and helps to manage the risks of investment tools for long positioning and present hedging opportunity skills of investment tools in portfolio diversifications. With the help of DCC-FIGARCH-t model, both long memory properties in volatility and time rates volatility spillover structure were explored.

When the results obtained using the TVP VAR model are examined, it is understood that the dynamic linkage between ECO, WTI and PSE that changes over time is weak and this relationship becomes stronger in periods such as Covid-19, European sovereign debt crisis (2012), Iran sanctions (2012) Lebanon Israel crisis (2006) and Paris climate agreement (2016). On the other hand, volatility clusters were found in crude oil and alternative energy and technology returns. For

Table 7: DCC-FIGARCH-t (1, d, 1) Model Results.

	FIGARCH Model			
Panel 1: Estimates of the univariate	OIL	PSE	ECO	
	0.000249**	0.0003566***	0.000185*	
Consi. (m)	-0.00012545	-0.00002972	-0.00010574	
Correct (v)	0.009216***	1.032569***	2.481320***	
Const. (V)	-0.0024461	-0.30248	-0.78988	
	0.965349***	0.532241***	0.424027***	
d-FIGARCH	-0.05295	-0.11878	-0.057908	
<i>.</i>	0.071314	0.100497	0.146596**	
Ø _{Arch(1)}	-0.056066	-0.065098	-0.068658	
0	0.896940***	0.519522***	0.482010***	
P _{Garch(1)}	-0.015627	-0.13779	-0.094093	
anel 2: Estimates of the Multivariate		Model		
- 1- 1	0.025814***			
aipna	-0.0050979			
h - t-	0.964754***			
Deta	-0.0087951			
<i>٤</i> ۲	7.642523***			
u -	-0.43013			
Log L		47733.5		
rho oil pse		0.2169		
`rho oil eco		0.2997		
rho pse eco		0.7055		
anel 3: Diagnostic tests				
-	15.9217	32.2787**	43.2901***	
Qs (20)	[0.7214684]	[0.0404173]	[0.0018713]	
2 (22)	9.28357	23.8057	21.7644	
Qs (20)	[0 9793839]	[0 2509822]	[0 3534305]	



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Figure 5: Time-Varying conditional correlation between variables.



Figure 6: Time-Varying Hedge Ratio Computed from DCC-FIGARCH-t model.

Table 8.	TV-	Hedge	Ratio	Summary	Statistics
I able o.	. I V -	neuge	Natio	Summary	Statistics.

	Mean	Min	Max	Std Dev
ECO/OIL	0.27084	-0.15267	1.0261	0.17595
OIL/ECO	0.38273	-0.39346	2.4673	0.25525
PSE/OIL	0.12605	-0.20609	0.6448	0.12689
OIL/PSE	0.42814	-1.0744	3.0319	0.41139
PSE/ECO	0.43406	0.1247	1.4124	0.12828
ECO/PSE	1.2014	0.29234	2.247	0.26551
Long/Short represents that the first asset is long and the second asset is shorted in a portfolio				





Figure 7: Time Varying Optimal Portfolio Weights.

rable 9: Fortiono weights.					
	Min	Mean	Max	Std. dev.	
ECO/PSE	0	0.042	1	0.12496	
PSE/ECO	0	0.958	1	0.12496	
OIL/ECO	0	0.408	1	0.22929	
ECO/OIL	0	0.592	1	0.22929	
OIL/PSE	0	0.17	0.79835	0.12659	
PSE/OIL	0.20165	0.83	1	0.12659	

T.L. O. D. (C.1' W. ' 1.

this reason, useful information shocks reach to all 3 assets and being eliminated at hyperbolic speed, also the volatility spillover lasted for a long time. These were found in DCC-FIGARCH-t model results. Detailed results are shown in Figure 3. After the 2008 global financial crisis, an increase was observed in the conditional correlation between investment tools. The result achieved here revealed that the technology sector could not contribute to hedging the risks caused by the long positioning in any of the selected investment tools. Time-varying hedge rates were considered to manage risks of 1\$ alternate energy long term position, WTI shorting of 0.27 \$ is needed. Moreover, to manage risks of 1\$ technology longterm position, WTI shorting of 0.13 \$ is needed. Especially to manage risks of 1\$ investment in technology category, 0.43 \$ investment should be made in the alternative energy category. Considering the hedging opportunities, technology category cannot offer serious opportunities in comparison to other investment alternatives and the main reason is the high correlation in alternative energy category.

DCC-FIGARCH-t model used in this study creates binary portfolios with investment tools by making use of the conditional variance and covariance matrices. The average weight of ECO/OIL assets in the study is 0.59. Given this result, it can be concluded that a portfolio of 1\$ should consist of 0.59 \$ clean energy and 0.41\$ WTI futures. According to the results of this study, correlation between clean energy (ECO) and technology (PSE) should be 75% and it should be noted that the technology sector cannot offer any hedging opportunities since it is relatively high. Hedging the longterm positioning risks in alternative energy and technology category by using the short-term positioning investments should be made in WTI future. On the other hand, long positioning risks of WTI futures can be compensated by clean energy asymmetric positions. Furthermore, it was observed that technology industry also offers similar hedging opportunities. For the investors who do not make portfolio diversification between two highly correlated investment instruments such as alternative energy and technology, it can be recommended to involve WTI futures in the portfolio, which will offer serious opportunities for managing the risks.

In this paper, it is aimed to fill a remarkable gap in literature by modeling the volatility of financial assets by using a more robust method with the DCC-FIGARCH-t model. Sadorsky [13] previously listed the short memory models such as Dynamic conditional correlation, Constant Conditional Correlation, etc. Although this subject has been examined by using models, the present paper is the first study examining these relationships by using models taking into account that information shocks that affect financial assets disappear at



hyperbolic speed. This is the point, which differentiates it from the previous studies.

While S and P Global Clean Energy Index, The MSCI Global Alternative Energy Index, MSCI World Information Technology index, and many other similar indices have been used in previous studies, but the energy and technology category indices haven't been included. This causes the most important constraints. By including more energy and technology indices in future studies, it will also be possible to make a selection between multiple models by the predictive performance. Considering the multivariate Fractional GARCH models, which take into account that time series are fractal (self-similarity), instead of short memory (CCC, BEKK, DCC GARCH, etc.) models, which have been used many times in modeling the return volatility of renewable energy and technology sectors, will offer important advantages to investors in terms of portfolio optimization and hedging opportunities.

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